

CSCE
478/878

Lecture 9:
Hidden
Markov
Models

Stephen Scott

Introduction

Outline

Markov
Chains

Hidden
Markov
Models

CSCE 478/878 Lecture 9: Hidden Markov Models

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- Useful for modeling/making predictions on *sequential data*
- E.g., biological sequences, text, series of sounds/spoken words
- Will return to *graphical models* that are *generative*

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- Markov chains
- Hidden Markov models (HMMs)
 - Formal definition
 - Finding most probable state path (Viterbi algorithm)
 - Forward and backward algorithms
- Specifying an HMM

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The Markov Property

Begin and End
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Discrimination

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- Focus on nucleotide sequences: Sequences of symbols from alphabet $\{A, C, G, T\}$
- The sequence “CG” (written “CpG”) tends to appear more frequently in some places than in others
- Such *CpG islands* are usually 10^2 – 10^3 bases long
- Questions:
 - 1 Given a short segment, is it from a CpG island?
 - 2 Given a long segment, where are its islands?

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- Model will be a CpG *generator*
- Want probability of next symbol to depend on current symbol
- Will use a standard (non-hidden) Markov model
 - Probabilistic state machine
 - Each state emits a symbol

Markov Chains

Modeling CpG Islands (cont'd)

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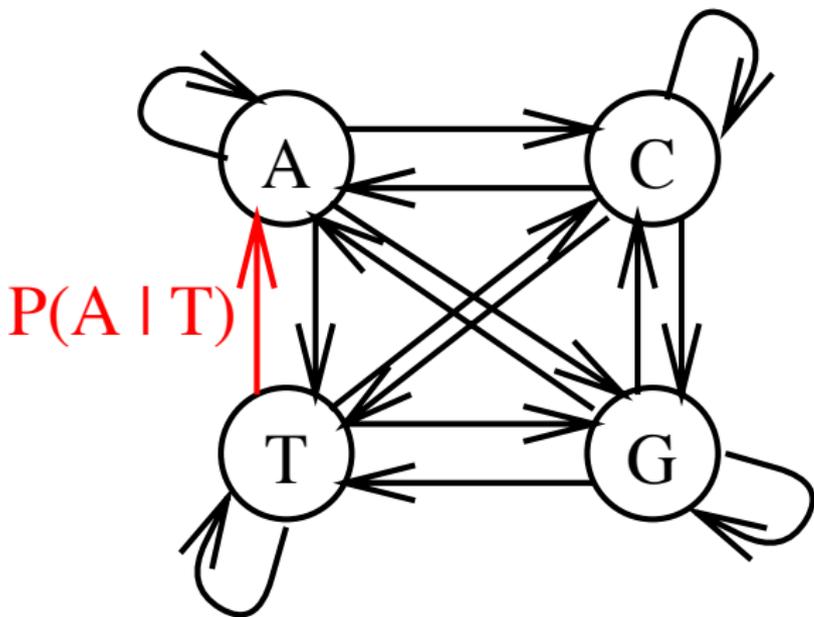
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- A *first-order* Markov model (what we study) has the property that observing symbol \mathbf{x}_i while in state π_i depends *only* on the previous state π_{i-1} (which generated \mathbf{x}_{i-1})
- Standard model has 1-1 correspondence between symbols and states, thus

$$P(\mathbf{x}_i \mid \mathbf{x}_{i-1}, \dots, \mathbf{x}_1) = P(\mathbf{x}_i \mid \mathbf{x}_{i-1})$$

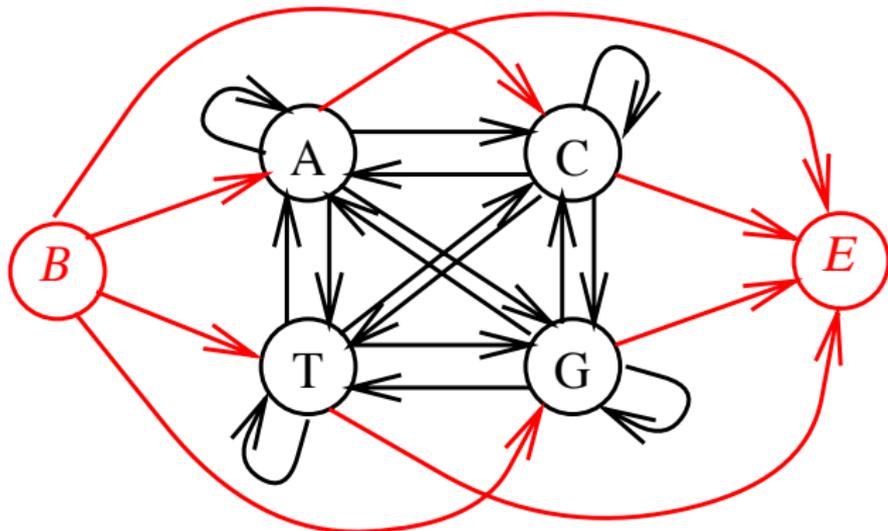
and

$$P(\mathbf{x}_1, \dots, \mathbf{x}_L) = P(\mathbf{x}_1) \prod_{i=2}^L P(\mathbf{x}_i \mid \mathbf{x}_{i-1})$$

Markov Chains

Begin and End States

- For convenience, can add special “begin” (B) and “end” (E) states to clarify equations and define a distribution over sequence lengths
- Emit empty (null) symbols x_0 and x_{L+1} to mark ends of sequence



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- How do we use this to differentiate islands from non-islands?
- Define two Markov models: islands (“+”) and non-islands (“-”)
 - Each model gets 4 states (A, C, G, T)
 - Take training set of known islands and non-islands
 - Let c_{st}^+ = number of times symbol t followed symbol s in an island:

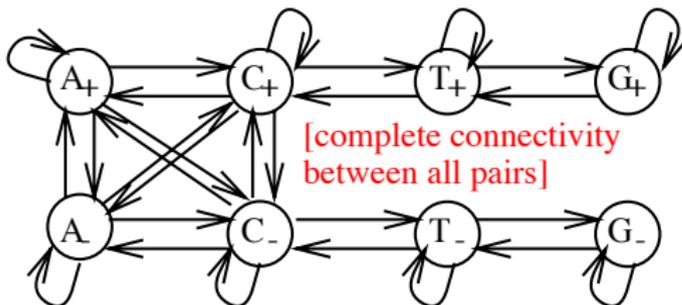
$$\hat{P}^+(t | s) = \frac{c_{st}^+}{\sum_{t'} c_{st'}^+}$$

- Now **score** a sequence $X = \langle \mathbf{x}_1, \dots, \mathbf{x}_L \rangle$ by summing the **log-odds ratios**:

$$\log \left(\frac{\hat{P}(X | +)}{\hat{P}(X | -)} \right) = \sum_{i=1}^{L+1} \log \left(\frac{\hat{P}^+(\mathbf{x}_i | \mathbf{x}_{i-1})}{\hat{P}^-(\mathbf{x}_i | \mathbf{x}_{i-1})} \right)$$

Second CpG question: Given a long sequence, where are its islands?

- Could use tools just presented by passing a fixed-width window over the sequence and computing scores
- Trouble if islands' lengths vary
- Prefer single, unified model for islands vs. non-islands



- Within the + group, transition probabilities similar to those for the separate + model, but there is a small chance of switching to a state in the - group

Hidden Markov Models

What's Hidden?

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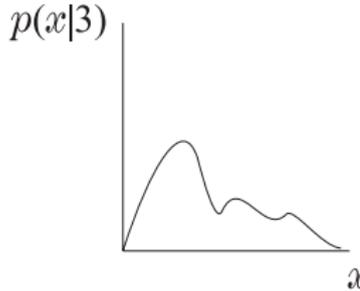
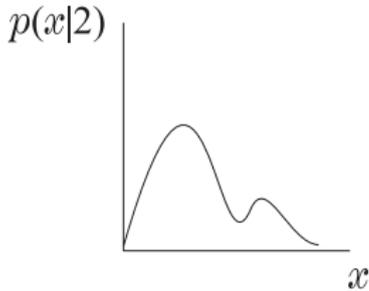
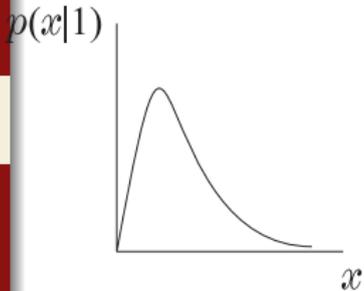
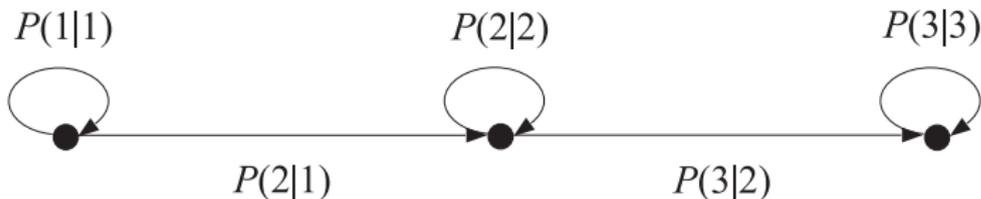
Baum-Welch

Structure

- No longer have one-to-one correspondence between states and emitted characters
 - E.g., was C emitted by C_+ or C_- ?
- Must differentiate the *symbol* sequence X from the *state* sequence $\pi = \langle \pi_1, \dots, \pi_L \rangle$
 - State transition probabilities same as before:
 $P(\pi_i = \ell \mid \pi_{i-1} = j)$ (i.e., $P(\ell \mid j)$)
 - Now each state has a prob. of emitting any value:
 $P(\mathbf{x}_i = \mathbf{x} \mid \pi_i = j)$ (i.e., $P(\mathbf{x} \mid j)$)

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What's Hidden? (cont'd)

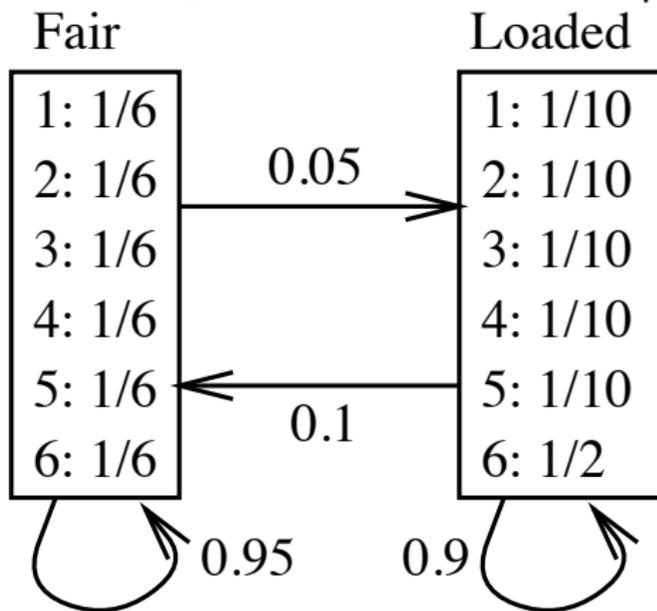


[In CpG HMM, emission probs discrete and = 0 or 1]

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Example: The Occasionally Dishonest Casino

Assume casino is typically fair, but with prob. 0.05 it switches to loaded die, and switches back with prob. 0.1



Given a sequence of rolls, what's hidden?

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The Viterbi Algorithm

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Structure

- Probability of seeing symbol sequence X and state sequence π is

$$P(X, \pi) = P(\pi_1 | 0) \prod_{i=1}^L P(\mathbf{x}_i | \pi_i) P(\pi_{i+1} | \pi_i)$$

- Can use this to find most likely path:

$$\pi^* = \operatorname{argmax}_{\pi} P(X, \pi)$$

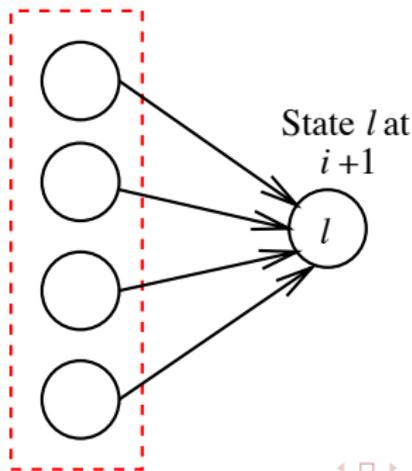
and trace it to identify islands (paths through “+” states)

- There are an exponential number of paths through chain, so how do we find the most likely one?

- Assume that we know (for all k) $v_k(i) =$ probability of most likely path ending in state k with observation \mathbf{x}_i
- Then

$$v_\ell(i+1) = P(\mathbf{x}_{i+1} | \ell) \max_k \{v_k(i) P(\ell | k)\}$$

All states at i



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Structure

Given the formula, can fill in table with dynamic programming:

- $v_0(0) = 1, v_k(0) = 0$ for $k > 0$
- For $i = 1$ to L ; for $\ell = 1$ to M (# states)
 - $v_\ell(i) = P(\mathbf{x}_i | \ell) \max_k \{v_k(i-1)P(\ell | k)\}$
 - $\text{ptr}_i(\ell) = \text{argmax}_k \{v_k(i-1)P(\ell | k)\}$
- $P(X, \pi^*) = \max_k \{v_k(L)P(0 | k)\}$
- $\pi_L^* = \text{argmax}_k \{v_k(L)P(0 | k)\}$
- For $i = L$ to 1
 - $\pi_{i-1}^* = \text{ptr}_i(\pi_i^*)$

To avoid underflow, use $\log(v_\ell(i))$ and add

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The Forward Algorithm

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Structure

Given a sequence X , find $P(X) = \sum_{\pi} P(X, \pi)$

Use dynamic programming like Viterbi, replacing max with sum, and $v_k(i)$ with $f_k(i) = P(\mathbf{x}_1, \dots, \mathbf{x}_i, \pi_i = k)$ (= prob. of observed sequence through \mathbf{x}_i , stopping in state k)

- $f_0(0) = 1, f_k(0) = 0$ for $k > 0$
- For $i = 1$ to L ; for $\ell = 1$ to M (# states)
 - $f_{\ell}(i) = P(\mathbf{x}_i | \ell) \sum_k f_k(i-1)P(\ell | k)$
- $P(X) = \sum_k f_k(L)P(0 | k)$

To avoid underflow, can again use logs, though exactness of results compromised

Given a sequence X , find the probability that \mathbf{x}_i was emitted by state k , i.e.,

$$\begin{aligned}
 P(\pi_i = k \mid X) &= \frac{P(\pi_i = k, X)}{P(X)} \\
 &= \frac{\overbrace{P(\mathbf{x}_1, \dots, \mathbf{x}_i, \pi_i = k)}^{f_k(i)} \overbrace{P(\mathbf{x}_{i+1}, \dots, \mathbf{x}_L \mid \pi_i = k)}^{b_k(i)}}{P(X)} \\
 &\qquad\qquad\qquad \underbrace{\hspace{10em}}_{\text{computed by forward alg}}
 \end{aligned}$$

Algorithm:

- $b_k(L) = P(0 \mid k)$ for all k
- For $i = L - 1$ to 1 ; for $k = 1$ to M (# states)
 - $b_k(i) = \sum_{\ell} P(\ell \mid k) P(\mathbf{x}_{i+1} \mid \ell) b_{\ell}(i + 1)$

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Structure

- Define $g(k) = 1$ if $k \in \{A_+, C_+, G_+, T_+\}$ and 0 otherwise
- Then $G(i | X) = \sum_k P(\pi_i = k | X) g(k) =$ probability that x_i is in an island
- For each state k , compute $P(\pi_i = k | X)$ with forward/backward algorithm
- Technique applicable to any HMM where set of states is partitioned into classes
 - Use to label individual parts of a sequence

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Specifying an HMM

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Structure

- Two problems: defining *structure* (set of states) and *parameters* (transition and emission probabilities)
- Start with latter problem, i.e., given a training set X_1, \dots, X_N of independently generated sequences, learn a good set of parameters θ
- Goal is to maximize the (log) likelihood of seeing the training set given that θ is the set of parameters for the HMM generating them:

$$\sum_{j=1}^N \log(P(X_j; \theta))$$

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Structure

- Estimating parameters when e.g., islands already identified in training set
- Let $A_{k\ell}$ = number of $k \rightarrow \ell$ transitions and $E_k(b)$ = number of emissions of b in state k

$$P(\ell | k) = A_{k\ell} / \left(\sum_{\ell'} A_{k\ell'} \right)$$

$$P(b | k) = E_k(b) / \left(\sum_{b'} E_k(b') \right)$$

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Structure

Be careful if little training data available

- E.g., an unused state k will have undefined parameters
- Workaround: Add *pseudocounts* r_{kl} to A_{kl} and $r_k(b)$ to $E_k(b)$ that reflect prior biases about probabilities
- Increased training data decreases prior's influence

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Structure

- Used for estimating params when state seq unknown
- Special case of *expectation maximization* (EM)
- Start with arbitrary $P(\ell | k)$ and $P(b | k)$, and use to estimate $A_{k\ell}$ and $E_k(b)$ as *expected* number of occurrences given the training set¹:

$$A_{k\ell} = \sum_{j=1}^N \frac{1}{P(\mathbf{X}_j)} \sum_{i=1}^L f_k^j(i) P(\ell | k) P(\mathbf{x}_{i+1}^j | \ell) b_\ell^j(i+1)$$

(Prob. of transition from k to ℓ at position i of sequence j , summed over all positions of all sequences)

¹Superscript j corresponds to j th train example 

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Specifying an HMM: The Baum-Welch Algorithm (cont'd)

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$$E_k(b) = \sum_{j=1}^N \sum_{i: x_i^j = b} P(\pi_i = k | X_j) = \sum_{j=1}^N \frac{1}{P(X_j)} \sum_{i: x_i^j = b} f_k^j(i) b_k^j(i)$$

- Use these (& pseudocounts) to recompute $P(\ell | k)$ and $P(b | k)$
- After each iteration, compute log likelihood and halt if no improvement

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Structure

- How to specify HMM states and connections?
- States come from background knowledge on problem, e.g., size-4 alphabet, $+/-$, \Rightarrow 8 states
- Connections:
 - Tempting to specify complete connectivity and let Baum-Welch sort it out
 - *Problem*: Huge number of parameters could lead to local max
 - Better to use background knowledge to invalidate some connections by initializing $P(\ell | k) = 0$
 - Baum-Welch will respect this

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Specifying an HMM: Silent States

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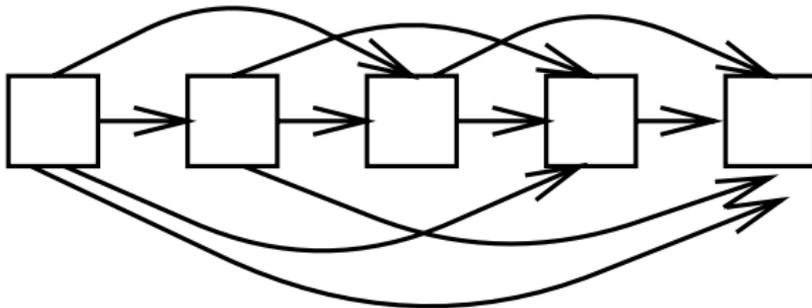
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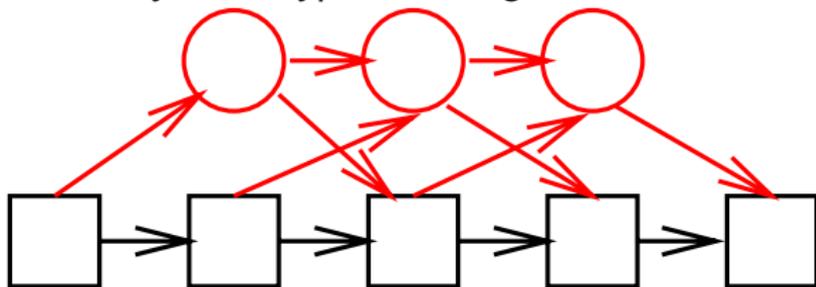
Structure

- May want to allow model to generate sequences with certain parts *deleted*
 - E.g., when aligning DNA or protein sequences against a fixed model or matching a sequence of spoken words against a fixed model, some parts of the input might be omitted



- Problem: Huge number of connections, slow training, local maxima

- *Silent states* (like begin and end states) don't emit symbols, so they can "bypass" a regular state



- If there are no purely silent loops, can update Viterbi, forward, and backward algorithms to work with silent states
- Used extensively in *profile HMMs* for modeling sequences of protein families (aka *multiple alignments*)